

On the Testing of Ground-Motion Prediction Equations against Small-Magnitude Data

by Céline Beauval, Hilal Tasan, Aurore Laurendeau, Elise Delavaud,
Fabrice Cotton, Philippe Guéguen, and Nicolas Kuehn

Abstract Ground-motion prediction equations (GMPE) are essential in probabilistic seismic hazard studies for estimating the ground motions generated by the seismic sources. In low-seismicity regions, only weak motions are available during the lifetime of accelerometric networks, and the equations selected for the probabilistic studies are usually models established from foreign data. Although most GMPEs have been developed for magnitudes 5 and above, the minimum magnitude often used in probabilistic studies in low-seismicity regions is smaller. Disaggregations have shown that, at return periods of engineering interest, magnitudes less than 5 may be contributing to the hazard. This paper presents the testing of several GMPEs selected in current international and national probabilistic projects against weak motions recorded in France (191 recordings with source–site distances up to 300 km, $3.8 \leq M_w \leq 4.5$). The method is based on the log-likelihood value proposed by [Scherbaum *et al.* \(2009\)](#). The best-fitting models (approximately $2.5 \leq LLH \leq 3.5$) over the whole frequency range are the [Cauzzi and Faccioli \(2008\)](#), [Akkar and Bommer \(2010\)](#), and [Abrahamson and Silva \(2008\)](#) models. No significant regional variation of ground motions is highlighted, and the magnitude scaling could be the predominant factor in the control of ground-motion amplitudes. Furthermore, we take advantage of a rich Japanese dataset to run tests on randomly selected low-magnitude subsets, and confirm that a dataset of ~ 190 observations, the same size as the French dataset, is large enough to obtain stable LLH estimates. Additionally we perform the tests against larger magnitudes (5–7) from the Japanese dataset. The ranking of models is partially modified, indicating a magnitude scaling effect for some of the models, and showing that extrapolating testing results obtained from low-magnitude ranges to higher magnitude ranges is not straightforward.

Introduction

A probabilistic seismic hazard analysis (PSHA) is usually carried out to establish a national seismic building code. Such analysis relies on the identification of the seismic sources (size, location, and occurrence probability) and the estimation of their capacity to produce ground motions. A ground-motion prediction equation (GMPE) is necessary for estimating the conditional probability that, if the earthquake occurs, a given acceleration threshold can be exceeded at the site of interest. The minimum magnitude used in PSHA studies varies from 5 down to 4, and even lower values in very low-seismicity regions such as Scandinavian countries. Disaggregation studies show that the whole range of magnitudes considered contributes to the PSHA, with the barycenter of the contributions depending on the return period considered and the seismicity level of the region. [Beauval *et al.* \(2006, 2008\)](#) showed that, in the active parts of France, magnitudes

contributing the most for return periods of 475 years are in the interval [5–5.5], but magnitudes in the 4–5 range are also responsible for a nonnegligible contribution to the hazard, even for return periods as large as 10,000 years. Therefore, there is a need for reliable predictions of ground-motion amplitudes over the whole magnitude range.

To develop a robust GMPE, a large accelerometric dataset containing a wide range of magnitudes and source-to-site distances is required. In low-seismicity regions such as France, the bulk of the data consists of low-magnitude recordings ($M_w < 4.5$). Several studies showed that, due to magnitude-scaling problems, equations based on low-magnitude datasets are not able to correctly predict the ground motions of moderate-to-large magnitudes ($M_w \geq 5$; see [Youngs *et al.*, 1995](#); [Bommer *et al.*, 2007](#); [Cotton *et al.*, 2008](#)). The solution proposed up to now is to select the GMPEs among published

equations established from strong motions recorded in higher seismicity regions (either global or region-specific models; Bommer *et al.*, 2010). The large majority of these equations have been developed for magnitude 5 and above (Douglas, 2011). This is the case for all GMPEs selected within the Seismic Hazard Harmonization in Europe (SHARE) project for application in crustal active regions (see Delavaud, Cotton, *et al.*, 2012). The main objective of the SHARE project is to provide a seismic hazard model for the Euro-Mediterranean region. In PSHA studies led in low-seismicity regions, the use of models based on allogenous data and developed for $M_w \geq 5$ rely on two assumptions: (1) for the same magnitude and distance, the ground motions produced do not vary much from one shallow active region to the other; (2) the models established from moderate-to-large magnitudes have the ability to predict amplitudes of motions produced by lower magnitudes.

The first assumption, the regional variability of ground motions, is currently a strongly debated issue, mainly due to the current lack of data. The issue will be completely solved only when local models are available in all crustal regions; however, this will not happen in the near future (Stafford *et al.*, 2007). Some authors believe that ground motions do not vary much regionally, at least for moderate-to-large magnitudes, as long as the same tectonic environment is considered (Bommer, 2006; Stafford *et al.*, 2007). In fact, all global models based on a database including data from different regions of the world assume that ground motions are not regionally dependent (e.g., The Next Generation Attenuation [NGA] models; Abrahamson *et al.*, 2008). On the opposite side, some authors have highlighted significant regional dependence (e.g., Luzi *et al.*, 2006, for moderate magnitudes in Italy); however, it is often based on restricted regional datasets. For low magnitudes, different recent publications show statistically significant regional differences in ground motions (e.g., Bakun and Scotti, 2006, based on macroseismic intensities; Atkinson and Morrison, 2009; Chiou *et al.*, 2010), while others did not evidence such a discrepancy (Douglas, 2004). Atkinson and Morrison (2009) and Chiou *et al.* (2010) found that ground-motion amplitudes from small earthquakes in northern California are lower on average than those for southern California. They observed that these differences are no longer significant for magnitudes larger than ~ 6 . The ground motions from smaller earthquakes may be more sensitive to differences in crustal structure or crustal stress states. Because stress drops of small earthquakes appear to be magnitude-dependent, the regional dependence of the average stress drop could result in different ground motions (Chiou *et al.*, 2010). Besides, as stress is dependent on depth, the focal depth of small-magnitude earthquakes might also play a significant role in the scaling of ground motions.

The second assumption resides in the way the models account for the scaling of ground motions with magnitude. Following the selection criteria of Bommer *et al.* (2010), a model should include nonlinear scaling of ground-motion amplitudes with magnitude. If the w^{-2} source model is

assumed, the corner period of the spectrum varies with the magnitude, and the scaling law of the source spectrum amplitude becomes a nonlinear function of magnitude (Fukushima, 1996). The scaling relation of the spectrum amplitude with M_w can be approximated as a quadratic function, and the coefficient of the M^2 term should be negative, implying that the rate of decrease in spectral amplitude with decreasing magnitude is accentuated (e.g., Zhao *et al.*, 2006; Bindi *et al.*, 2009). However, these constraints on the magnitude scaling do not allow for the extrapolation of GMPEs at the limits or beyond their range of applicability (Bommer *et al.*, 2007). This limitation might be attributed to different effects: the magnitude scaling of ground motions that decreases with increasing magnitude, the stronger decay of small-magnitude motions with respect to larger-magnitude motions, and/or the scaling of the stress drop with magnitude (Cotton *et al.*, 2008; Atkinson and Boore, 2011). In the NGA models, authors proposed more complex magnitude scaling (e.g., bilinear and trilinear), but still these equations derived from larger-magnitude earthquake recordings can overestimate the ground motions produced by smaller-magnitude events (Atkinson and Morrison, 2009). Bommer *et al.* (2007) concluded that for modeling the magnitude scaling over an extended range, this scaling could be linear at low magnitudes and then allow for a quadratic fall-off in slope in the upper range. They are currently working to understand which approach would be the best. Chiou *et al.* (2010) and Atkinson and Boore (2011) provided an update of their NGA models extended to lower magnitudes (down to 3.0) based on data from western North America (median amplitudes are updated for three frequencies). It is not known yet if these modified models are predicting correctly low-magnitude ground motions elsewhere.

Purposes of the Study

In the short term, except for these two models previously mentioned in this paper, reliable equations for predicting ground-motion amplitudes from very low to large magnitudes do not exist. Therefore, the purpose of the present study is to test the models selected in the SHARE project on the low-magnitude dataset available for the French Accelerometric Network. The first aim is to evaluate how these models perform on low-magnitude ground motions (M 4–5) in a magnitude range that contributes to PSHA estimates obtained in this region. The second aim is to analyze the results in terms of regional variability of ground motions, keeping in mind that the interpretation will be limited as both aforementioned problems (magnitude scaling and regional variation) play a role and cannot be analyzed separately.

The first testing of GMPEs against weak motions recorded in active regions of Western Europe, using a reproducible technique, was proposed by Scherbaum *et al.* (2004) and applied in two subsequent studies (Drouet *et al.*, 2007; Hintersberger *et al.*, 2007). Scherbaum *et al.* (2004) proposed a likelihood-based technique to rank models according to their fit to observed data, through a categorization scheme. They

illustrated the method on the records of a unique event, the Saint Dié earthquake (22 February 2003, M_w 4.5 according to [Drouet *et al.*, 2010](#)), at 13 rock stations. [Hintersberger *et al.* \(2007\)](#) performed an update of this study, using the same method applied on acceleration records from five earthquakes (M_w 3.6–5.1) in the border region of Germany, France, and Switzerland, resulting in a dataset of 61 records with distances up to 300 km. [Drouet *et al.* \(2007\)](#) used 15 accelerometric records obtained in the Pyrenees from three earthquakes (M_w 3.7, 3.7, and 3.9). In these studies, roughly the same set of GMPEs was tested, recordings at rock stations were considered, and all frequencies available were mixed in the testing. The resulting best-fitting GMPEs were models derived from different tectonic contexts: Europe, the western United States, and Japan. Therefore, these previous studies did not highlight regional variations of ground motions, but the results were interpreted with great caution, as the datasets were quite restricted. They posed the problem of testing GMPEs derived from earthquakes with larger magnitudes than available in the datasets. The current study includes the dataset from [Drouet *et al.* \(2007\)](#), as well as part of the accelerometric data used in [Scherbaum *et al.* \(2004\)](#) and [Hintersberger *et al.* \(2007\)](#). GMPEs have evolved a lot, and in the following none of the GMPEs tested in these previous studies will be used as they all have been superseded.

The method chosen here for quantifying the goodness-of-fit of a GMPE to a dataset is the [Scherbaum *et al.* \(2009\)](#) technique (detailed in [Method for Testing GMPEs against Observations](#)). This method is new and has not been widely applied yet, and hence we propose synthetic tests to evaluate the meaning of the LLH value reflecting the fit between a model and the data. Another issue that is not clear yet is the minimum number of data required for the [Scherbaum *et al.* \(2009\)](#) method to yield stable results. When applying the technique on a given dataset, we have no argument for stating that the results are independent of the sample or can be considered as representative of the region in this magnitude range. Therefore, in the last part of the paper, we take advantage of a Japanese dataset to propose an answer to this question. As the Japanese dataset extends over a wide magnitude range, the question of the magnitude scaling is also addressed by comparing results obtained from low and moderate-to-large magnitude datasets.

Method for Testing GMPEs against Observations

The recent method introduced by [Scherbaum *et al.* \(2009\)](#) is chosen for testing the models against the data. [Scherbaum *et al.* \(2009\)](#) provides a ranking criterion based on information theory (see the original paper for a detailed description of the theory behind it). This technique is based on the probability for an observed ground motion to be realized under the hypothesis that a model is true. It provides one value, the negative average log-likelihood LLH ([Delavaud *et al.*, 2009](#)), that reflects the fit between the data and model:

$$\text{LLH} = -\frac{1}{N} \sum_{i=1}^N \log_2(g(x_i)), \quad (1)$$

with N the number of observations x_i , and g the probability density function (PDF) predicted by the GMPE (normal distribution). The ranking of models according to their fit to the data is then straightforward. In theory, it can be applied to whatever amounts of data are available, but the results are expected to be more stable if the testing is performed on a large dataset. A small LLH indicates that the candidate model is close to the model that has generated the data, while a large LLH corresponds to a model that is less likely to have generated the data.

We propose to take advantage of synthetic data to understand the meaning of the LLH value (Fig. 1). In this respect, we step back from the information theory perspective and simply concentrate on the calculated LLH values. Synthetic datasets are generated from an original Gaussian distribution, and distributions with modified characteristics (in terms of mean and sigma) are tested against these synthetic datasets. Results are displayed with an increasing number of synthetic data (generated and tested), generating a new random dataset for each run so that the stability of the results can be verified. If testing the same distribution on the simulated dataset, mean LLH values obtained are close to 1.4–1.5. Then, if testing distributions that differ from the original one, mean LLH values are higher. For a distribution with a mean equal to the original mean plus one sigma, or a sigma twice the original sigma, LLH values are around 2.0 (both distributions are roughly providing the same fit to the data). In the worst case considered in this example, the tested distribution has a mean equal to the original mean plus 2.5 sigma, and a sigma equal to 0.8 times the original sigma, producing mean LLH values as high as 9–10. These synthetic tests give us an idea of the LLH value to expect, depending on the distribution of the observations with respect to the model. In addition, these simulations also yield a rough idea of the minimum number of observations required for obtaining a stable mean LLH. Based on the results in Figure 1, mean LLH values are reached from ~ 40 observations and higher. However, here the synthetic dataset is perfectly distributed according to a normal distribution, which is not true in the real cases. Hence, 40 observations must be considered as the minimum, and more tests must be performed on different observation datasets to clearly define this minimum number of data.

GMPEs Best Adapted to the French Weak-Motion Data

Description of the Data

The French Accelerometric Network (Réseau Accélérométrique Permanent, RAP; Fig. 2) has been operating for roughly 15 years, with more and more stations installed since 1996 ([Péquegnat *et al.*, 2008](#)). Only earthquakes recorded in metropolitan France will be taken into account here

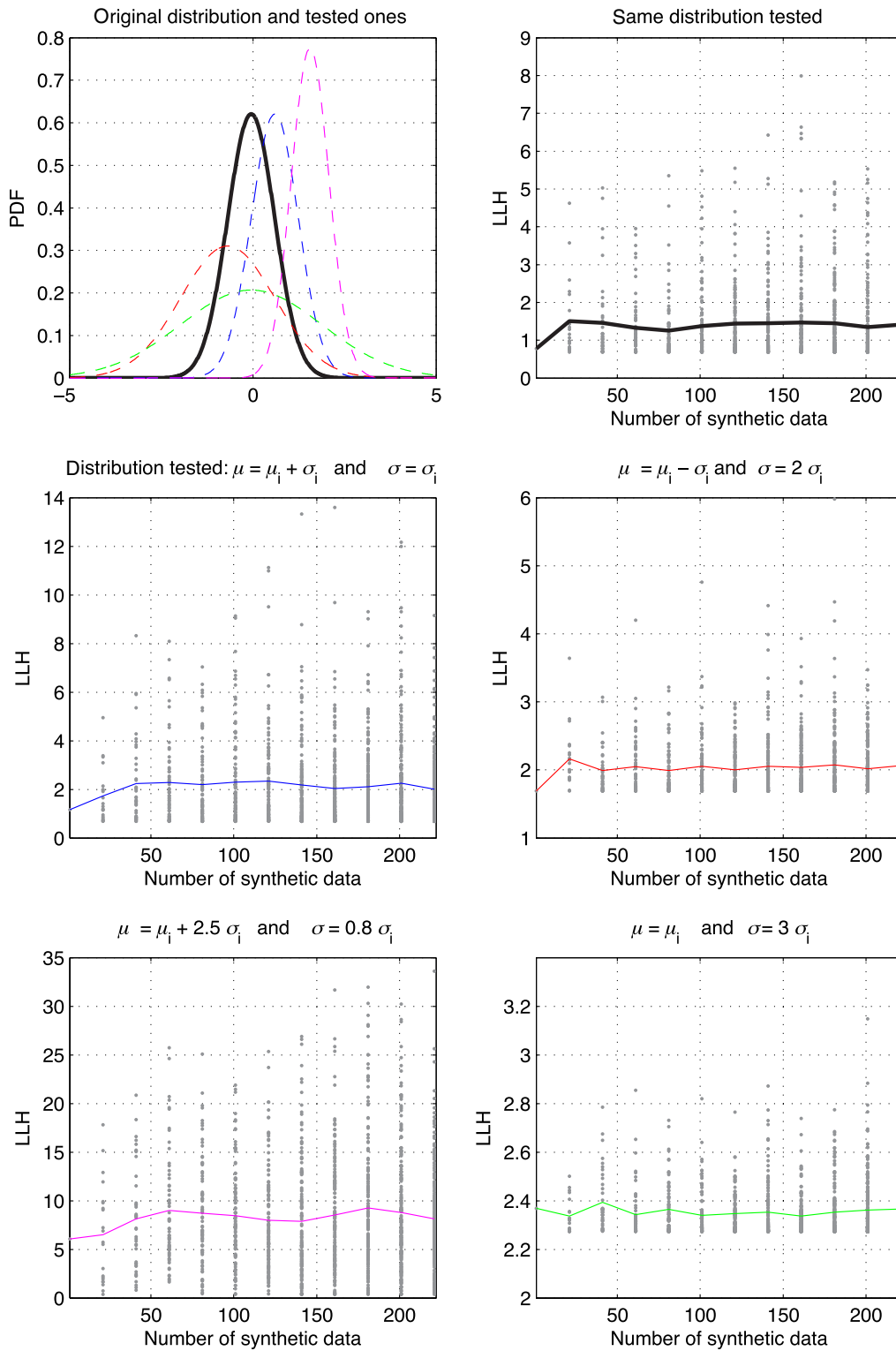


Figure 1. Synthetic data simulation to evaluate the meaning of LLH values. Upper left: the original distribution (black) and the distributions that are tested. The original distribution (μ_i, σ_i) corresponds to the ground motion predicted by the Akkar and Bommer (2010) model for a magnitude M_w 4 at $R_{JB} = 40$ km (natural logarithm of peak ground acceleration [PGA] in $m \cdot s^{-2}$). In the five other graphs, a candidate distribution is tested against a dataset generated from the original distribution. The mean and sigma of the candidate distribution are displayed in the title of each graph. The synthetic dataset is increased step by step (from 1 to 221 samples; the whole sample is randomly generated at each step). Individual LLH (gray points) and mean LLH values (solid line) are computed for each run. The color version of this figure is available only in the electronic edition.

(a separate study would be required for earthquakes belonging to the tectonic context of the French Antilles). The RAP stations are dial-up or continuous recording stations. They consist of one three-component broadband accelerometric sensor (kinematic episensors, except for some of the oldest stations having Guralp CMG5). They are connected to a 24-bit three-component digitizer sampling at 125 Hz. The full scale of the channel corresponds to $\pm 1g$ for all of the stations used in this paper. The useful frequency band is 0–50 Hz. Only offset correction is applied to the data without any additional filtering. The data used in this paper have been visually cleaned by checking the signal-to-noise ratio and the time accuracy on the three components.

Using only high-quality accelerograms, a total of 16 events with moment magnitudes between 3.8 and 4.5 are available (Figs. 2 and 3; Table 1). These earthquakes belong to the seismically active parts of France (Pyrenees, Alps, and Lower Rhine Embayment), which have been classified as active shallow crustal regions in [Delavaud, Cotton, et al. \(2012\)](#). Most of the GMPEs will be tested outside their validity range, but to limit the extrapolation below their minimum magnitude bound, moment magnitudes considered here are greater than or equal to M_w 3.8. Only a selection of the available M_w 3.8 events is included in this study (the events with the greatest number of recordings). Thus, GMPEs are tested against a dataset that is roughly homogeneously distributed in terms of magnitude, and the results are not influenced by the presence of more low magnitudes than larger ones (Fig. 3). The level of knowledge of the site conditions in the French Accelerometric Network varies greatly from one station to another. Sites are classified according to the four ground categories defined in the current [Eurocode 8](#)

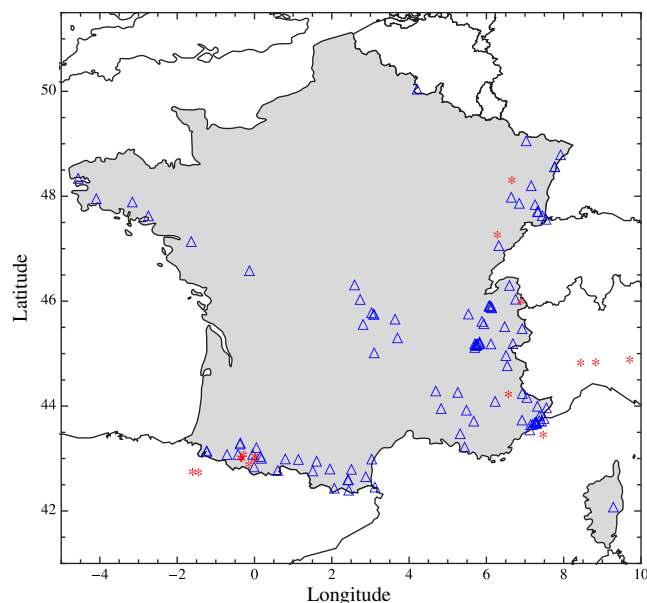


Figure 2. Location of earthquakes considered in the study (stars; see Table 1) and stations of the French Accelerometric Network RAP (triangles). The color version of this figure is available only in the electronic edition.

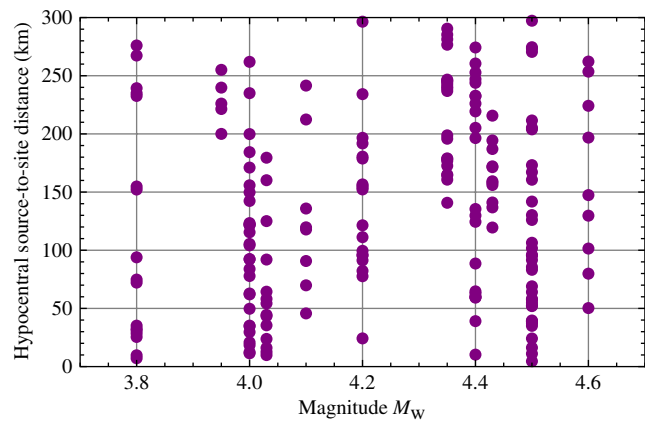


Figure 3. Distribution of the data used in this study, in terms of source-to-site distance and magnitude M_w (191 recordings in total; see Table 1). The color version of this figure is available only in the electronic edition.

(2004). At some sites, some geophysical and geological studies have been conducted. For the sites with a medium-to-high confidence level in the estimation of the shear-wave velocity in the top 30 m, the estimated V_{S30} has been used. Elsewhere, for the NGA models relying on V_{S30} , a mean V_{S30} value corresponding to the Eurocode 8 class has been used (1000, 600, 250, and 100 m/s for A, B, C, and D, respectively). We refer the reader to the [Régnier et al. \(2010\)](#) report for a detailed study on the site conditions of the accelerometric stations. Considering source–site distances of up to 300 km, a total of 191 recordings are obtained (Table 1). As the considered magnitudes are small, the size and extension of the fault planes have negligible impact on the calculation of the different distance measures. Moment magnitudes are available for 12 out of 16 events ([Drouet et al., 2010](#)). For the 4 remaining events, a magnitude correlation is applied ([Drouet et al., 2010](#)). Focal mechanisms are known for 12 out of the 16 events (Table 1). However, the influence of the style of faulting is unlikely to be significant for small events that can be approximated by point sources for most recording locations.

Description of the GMPEs Considered

Several recent models have been selected for testing the accelerometric dataset (Table 2; Fig. 4), although more than 180 equations are listed in [Douglas \(2011\)](#) for elastic response spectral ordinates. This short list consists of equations developed for active shallow crustal regions, in terms of moment magnitude and including a nonlinear magnitude scaling term (except for [Cauzzi and Faccioli, 2008](#)). The [Zhao et al. \(2006\)](#), [Cauzzi and Faccioli \(2008\)](#), [Chiou and Youngs \(2008\)](#), and [Akkar and Bommer \(2010\)](#) models have been selected within the SHARE project for application in shallow crustal regions (see [Delavaud, Cotton, et al., 2012](#), for a detailed description of the selection process). Here, two more NGA models often considered in current engineering seismology projects are tested, in particular, [Abrahamson](#)

Table 1
Description of the Earthquakes and Corresponding Recordings Used in This Study

Event Dates (dd/mm/yyyy)	M_w	Mech.*	Longitude (°)	Latitude (°)	Depth (km)	Number of Stations		Reference Coordinates	Reference Mechanism
						≤300 km	≤200 km		
31/10/1997	4	U	6.57	44.26	2	11	10	Drouet <i>et al.</i> (2010)	No reference
21/08/2000	4.4 [†]	U	8.44	44.86	10	11	11	RéNaSS	No reference
25/02/2001	4.5 [‡]	R	7.47	43.49	14	5	5	(BCSF) 2001 [§]	BCSF (2001) [§]
16/05/2002	4	N	-0.143	42.922	9.5	9	9	Drouet <i>et al.</i> (2010)	Chevrot <i>et al.</i> (2011)
11/02/2002	3.8	N	-0.33	43.04	5	5	5	Drouet <i>et al.</i> (2010)	Chevrot <i>et al.</i> (2011)
12/12/2002	4	N	-0.28	43.11	10	9	7	Drouet <i>et al.</i> (2010)	Chevrot <i>et al.</i> (2011)
21/01/2003	3.8	N	-0.36	43.05	10	12	7	Drouet <i>et al.</i> (2010)	Chevrot <i>et al.</i> (2011)
22/02/2003	4.5	N	6.66	48.34	10	13	9	Drouet <i>et al.</i> (2010)	BCSF (2003) [§]
11/04/2003	4.3 [†]	U	8.97	44.81	5	21	10	RéNaSS	No reference
23/02/2004	4.2	SS	6.28	47.3	10	19	17	Drouet <i>et al.</i> (2010)	BCSF (2004) [§]
18/09/2004	4.6	N	-1.6	42.78	2	9	6	Drouet <i>et al.</i> (2010)	Chevrot <i>et al.</i> (2011)
30/09/2004	4.1	N	-1.45	42.77	10	8	6	Drouet <i>et al.</i> (2010)	RAP (2005)
08/09/2005	4.4	SS	6.87	46.01	10	22	12	Drouet <i>et al.</i> (2010)	RAP (2005)
17/11/2006	4.5	N	0.01	43.08	9.7	18	15	Drouet <i>et al.</i> (2010)	Sylvander <i>et al.</i> (2008)
30/07/2007	4.0 [†]	U	9.71	44.92	10	5	0	RéNaSS	No reference
15/11/2007	4.0 [†]	N	0.0	43.01	8	14	14	BCSF (2008) [§]	BCSF (2008) [§]

Magnitudes have all been calculated by Drouet *et al.* (2010), except when specified.

*Mech., mechanism; SS, strike slip; N, normal; R, reverse; U, unknown; BCSF, Bureau Central Sismologique Français.

[†]Calculated from the RéNaSS local magnitude using the Drouet *et al.* (2010) M_w - $M_{L, RéNaSS}$ correlation (RéNaSS stands for the Réseau National de Surveillance Sismique).

[‡]Moment magnitude from ETH-Zurich.

[§]Bureau Central Sismologique Français.

and Silva (2008) and Boore and Atkinson (2008), as well as the new Chiou *et al.* (2010) equation extended to lower magnitudes. Furthermore, the Bindi *et al.* (2009) model, which showed to predict well the SHARE strong-motion dataset (Delavaud, Cotton, *et al.* 2012), is also considered.

Akkar and Bommer (2010) have developed a pan-European equation, predicting the geometrical mean of horizontal pseudospectral accelerations for magnitudes ranging from 5 to 7.6, at distances up to 100 km (Joyner and Boore distance). The spectral period range is 0.05–3 s. The generating dataset covers several countries in Europe and the Middle East, from moderate to high seismicity.

The Cauzzi and Faccioli (2008) model predicts geometrical mean of accelerations for magnitudes ranging from 5 to 7.2, at distances up to 150 km (hypocentral distance). The spectral period range is 0.05–20 s. Such long periods are crucial for the prediction of ground motions for bridges and tall buildings. This equation was initially developed for application in Italy, but it is based on a worldwide crustal earthquake dataset. A large part of this dataset (~80%) comes from the Japanese K-NET strong-motion network (see the Data and Resources section), and 5% comes from Europe and Turkey. The model handles two definitions for the site conditions: either directly using V_{S30} as the predictor variable or using the Eurocode 8 ground categories (Eurocode 8, 2004). This model has one limitation as it is defined for hypocentral distances larger than 15 km.

NGA models (Abrahamson *et al.*, 2008) have been developed from a worldwide dataset (including events from the

Euro-Mediterranean region) for predicting ground motions in the western United States, at distances up to 200 km and for spectral periods ranging from 0.01 to 10 s. Analytical models based on numerical simulations are included, providing constraints on the ground-motion scaling outside the range well constrained by the empirical data. Here, three of these models are considered. The Abrahamson and Silva (2008) and Boore and Atkinson (2008) equations apply for magnitudes greater than or equal to five. The Chiou and Youngs (2008) model is in theory applicable for magnitudes greater than or equal to four. These NGA models require some predictor variables that are not known for the French accelerometric stations and that must be estimated: depth-to-top of rupture and depth to the 1-km/s shear-wave velocity horizon. The Boore and Atkinson (2008) model uses the smallest number of predictor variables of the NGA models. All NGA models predict site response on the basis of the average shear-wave velocity in the top 30 m.

The Chiou *et al.* (2010) model has been developed for comparing weak ground motions between California and other active tectonic regions. For now, coefficients are available for three periods (peak ground acceleration [PGA], 1 s, and 0.3 s). The equation was developed for small-to-moderate shallow crustal earthquakes ($3 < M < 5.5$) for distances up to 200 km, and has been derived from Californian data. The specific goal of the authors is “to provide an empirical model that can be confidently used in the investigation of ground-motion difference between California and other active tectonic regions [...] where the bulk of ground-motion data from

Table 2
Ground-Motion Prediction Equations Used in This Study

GMPE Reference	GMPE Abbreviation	Magnitude Validity Bounds	Frequency Range (Hz)	Max. Source-Site Distance (km)*	Region of the Generating Dataset
Bindi et al. (2009)	B2009	4.0–6.9	0.5–33.33	100 (R_{JB})	Italy
Cauzzi and Faccioli (2008)	CF2008	5.0–7.2	0.05–20.0	150 (R_{HYP})	K-NET + worldwide
Kanno et al. (2006)	Ketal06	5.2–8.2	0.2–20.0	300 (R_{RUP})	Japan (depth < 30 km)
Chiou and Youngs (2008)	CY2008	4.0–8.0	0.1–100.0	200 (R_{RUP})	Worldwide
Chiou et al., 2010 (central)	Ccentral	3.0–5.5	PGA, 3.3, 1	200 (R_{RUP})	Central California
Chiou et al., 2010 (southern)	Csouth	3.0–5.5	PGA, 3.3, 1	200 (R_{RUP})	Southern California
Cotton et al. (2008)	Cetal08	4.1–7.3	0.3–100.0	100 (R_{RUP})	Japan
Zhao et al. (2006)	Z2006	5.0–7.3	0.2–20.0	300 (R_{RUP})	Japan
Boore and Atkinson (2008)	BA2008	5.0–8.0	0.1–100.0	200 (R_{JB})	Worldwide
Abrahamson and Silva (2008)	AS2008	5.0–8.5	0.1–100.0	200 (R_{RUP})	Worldwide
Akkar and Bommer (2010)	AB2010	5.0–7.6	0.33–20.0	100 (R_{JB})	Europe + Middle East

*JB, Joyner and Boore distance; HYP, hypocentral distance; RUP, rupture distance.

shallow crustal earthquakes is in the small-to-moderate magnitude range” ([Chiou et al., 2010](#), p. 1). Both the equations for southern California and central California will be tested.

[Bindi et al. \(2009\)](#) is an equation derived from Italian data only. The generating dataset is made up of magnitudes from 4.0 to 6.9 recorded at distances up to 100 km. The spectral period range is 0.03–2 s.

[Zhao et al. \(2006\)](#) is aimed at predicting ground motions in Japan. The dataset contains distances up to 300 km and magnitudes between 5 and 7.3 (crustal earthquakes). The spectral period range is 0.05–5 s. Most of the data have been recorded in Japan, except for a few overseas events providing short source distance recordings.

measure is different from one model to the other; for example, some models use the Joyner and Boore distance measure (which is measured horizontally on the surface; e.g., [Boore and Atkinson, 2008](#)), while others are based on the rupture distance (closest distance to the rupture plane). Each model is applied with its native distance measure. The [Beyer and Bommer \(2006\)](#) conversions are used to take into account different definitions in the horizontal component (geometrical mean, etc.). The NGA models predict site response on the basis of the average shear-wave velocity over the top 30 m (V_{S30}), whereas European equations take into account three generic site classes: rock, stiff soil, and soft soil (corresponding to shear-wave velocity intervals; [Eurocode, 2004](#)).

Parameter Compatibility Issue

All GMPEs considered in this study use the moment magnitude scale to characterize earthquake size. The distance

Results

The equations are tested against the homogeneous dataset described previously in this paper. Although most models

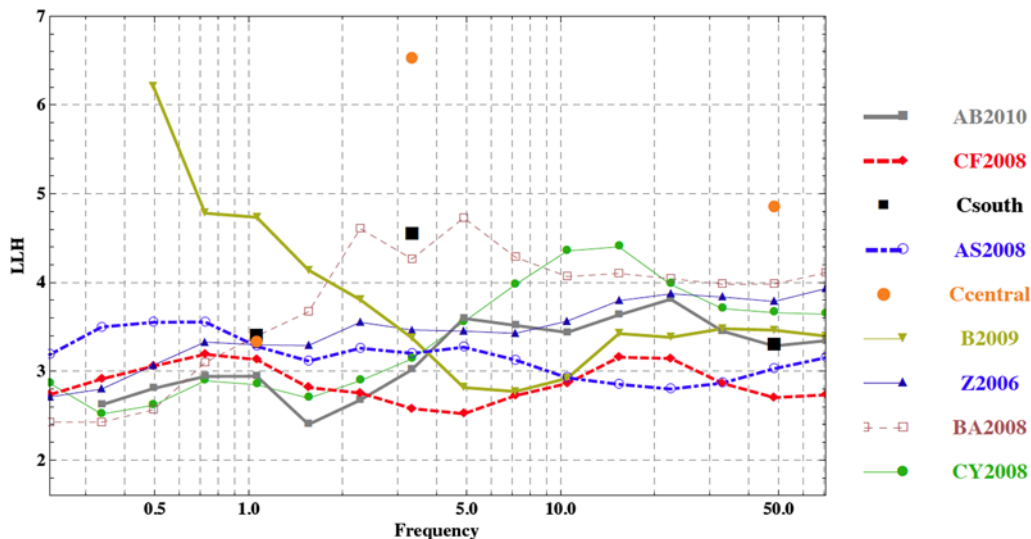


Figure 4. Quantifying the fit between observed spectral accelerations (French accelerometric data, Fig. 2) and corresponding predictions provided by a list of GMPEs; the LLH value versus frequency is shown. See Table 2 for the GMPE abbreviations. Note that [Chiou et al. \(2010\)](#); Ccentral and Csouth) is defined only for three frequencies (1, 3.33 Hz, and PGA); the results at PGA have been positioned at 50 Hz for graphical reasons. The color version of this figure is available only in the electronic edition.

have been developed for maximum distances varying from 100 to 200 km (Table 2), distances as far as 300 km are taken into account to ensure a minimum number of data (191 recordings). In a second step, the same calculations are performed on a reduced dataset selecting distances up to 200 km (143 recordings) to confirm that the results remain stable. As shown in Figure 4, all LLH values are roughly between 2.5 and 4.5. The synthetic tests showed that a perfect fit would yield $LLH = 1.4\text{--}1.5$ (Fig. 1, see [Method for Testing GMPEs against Observations](#)). Three models yield the lowest and most stable LLH values over the whole frequency range: [Cauzzi and Faccioli \(2008, CF2008\)](#), [Akkar and Bommer \(2010, AB2010\)](#), and [Abrahamson and Silva \(2008, AS2008\)](#). These three equations result in LLH values ranging from 2.5 to 3.5. The [Zhao et al. \(2006; Z2006\)](#) GMPE is not included in this best-fitting GMPE short list; compared to the three models mentioned previously in this paper, this GMPE yields slightly higher LLH on average over the whole frequency range. The [Abrahamson and Silva \(2008\)](#) model requires some parameters describing the site and the source, which are not well constrained (depth-to-top of rupture, fault dip, and down-dip rupture width) or not available (the depth to $V_S = 1000$ m/s). We performed sensitivity studies on these parameters, and we observed that, if using a reasonable (but still wide) range of values, there is very little impact (if any) on the LLH obtained. This might be due to the low magnitudes involved and the source–site distances available. Sensitivity tests performed for the [Chiou and Youngs \(2008\)](#) NGA model and the [Chiou et al. \(2010\)](#) model led to the same conclusions.

Some models show a good ability to predict the observations only for part of the frequency range. The [Chiou and Youngs \(2008\)](#) model performs roughly well only for low frequencies (< 3 Hz). Conversely, the [Bindi et al. \(2009\)](#) model predicts observations correctly only for higher frequencies, larger than ~ 3 Hz. These results highlight the need to test GMPEs as a function of spectral frequency. If mixing frequencies, a mean LLH value would be obtained that would not reflect correctly the goodness-of-fit of the model to the data. [Delavaud et al. \(2009\)](#) also observed a strong dependence with frequency while applying the LLH-based method on Californian data. It is interesting to observe that, if considering the results at 1 Hz, except for the [Bindi et al. \(2009\)](#) model, the LLH values are all concentrated in a narrow interval (2.9–3.3). For this frequency, the models' performances are comparable.

The [Chiou et al. \(2010\)](#) model, derived specifically for predicting ground motions for magnitudes less than 5.5, consists of two sets of coefficients each for central and southern California. The LLH values at 0.3 s and the PGA are very high using the coefficients for central California (6.5 at 0.3 s, and ~ 5 at the PGA). The equation for southern California fits the data well at 1 Hz and at the PGA, but it yields a higher LLH value than the rest of the models for the period 0.3 s. Low-magnitude ground motions in California might not be similar to low-magnitude ground motions in our target region.

However, this result must be taken with caution, as only three frequencies are available for comparison ([Chiou et al., 2010](#)).

Furthermore, to visualize the fit between the data and the predictions, a more classical technique is to display the residuals. The residual is the difference between the prediction and the observation in terms of the logarithm, normalized by the sigma of the model. Some residuals are displayed in Figures 5–8 to illustrate the fit for two of the best-fitting models (CF2008, Fig. 5; AB2010, Fig. 8) and for two of the models predicting higher LLH values (AB2008, Fig. 6; [Zhao et al., 2006; Z2006](#), Fig. 7). These histograms provide complementary insights on the fit between the models and data. They show that the observations are characterized by a higher variability than the predicted distributions, which is expected as motions from small earthquakes have proved to be more variable than motions from larger earthquakes (e.g., [Youngs et al., 1995](#)). The origin of this aleatory variability has not been identified yet (either a true physically based uncertainty or an uncertainty due to metadata; [Bommer et al., 2007](#)). Besides, whenever the median of observations does not fit the median of predictions, the models are overpredicting the amplitudes (e.g., Fig. 6 displays the results for the [Boore and Atkinson, 2008](#), model). This observation is also expected based on past studies (e.g., [Bommer et al., 2007](#)). Residual histograms are shown for varying maximum distances (300, 200, and 100 km). If we reduce the maximum distance to 200 km or consider only rock stations (and thus reduce the uncertainty on the site conditions), the ranking obtained for GMPEs remains stable. For 100 km, we believe that there is not enough data to derive reliable conclusions.

The equations that best fit the French accelerometric weak-motion dataset have been highlighted. Two models selected within the SHARE project for crustal regions fit the data reasonably well (CF2008; AB2010). No significant regional variation of ground motions is highlighted, and the magnitude scaling could be the predominant factor in the control of ground-motion amplitudes. However, clear explanations for the relative good performance of these models are not straightforward. These GMPEs are imported models, and they are applied at magnitudes lower than their minimum magnitude validity limits. The two other models selected in the SHARE project, [Zhao et al. \(2006\)](#) and [Boore and Atkinson \(2008\)](#), slightly overestimate the data, which is consistent with many recent studies (e.g., [Cotton et al., 2008](#)). One question that is naturally raised is whether the ranking deduced from these low-magnitude motions would hold if moderate magnitudes were available. However, at this stage, we have no argument to assert that, if available, stronger ground motions would also match these GMPEs. In the following, we make use of a Japanese dataset to tackle some of the unresolved questions that appeared during the testing on the weak-motion accelerometric dataset.

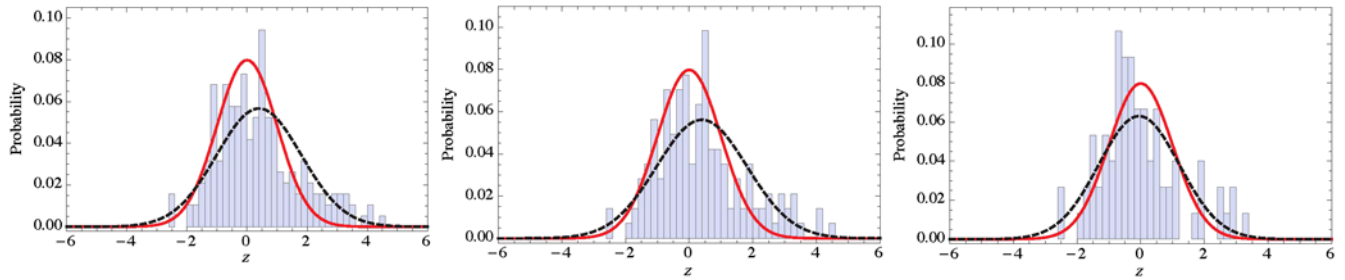


Figure 5. Histogram of the residuals superimposed on the standard normal distribution representing the [Cauzzi and Faccioli \(2008\)](#) model, using the French accelerometric subset described in the text, at 3.3 Hz. A residual z corresponds to $[\text{Log}(\text{observation}) - \text{Log}(\text{prediction})]/\sigma$. The Gaussian with mean and sigma calculated from the residuals is superimposed on the histogram (dashed curve). From left to right: the maximum distances considered are 300, 200, and 100 km (corresponding to $\text{LLH} = 2.58, 2.60, 2.18$). The numbers of data decrease accordingly: 191, 143, and 75 recordings. The color version of this figure is available only in the electronic edition.

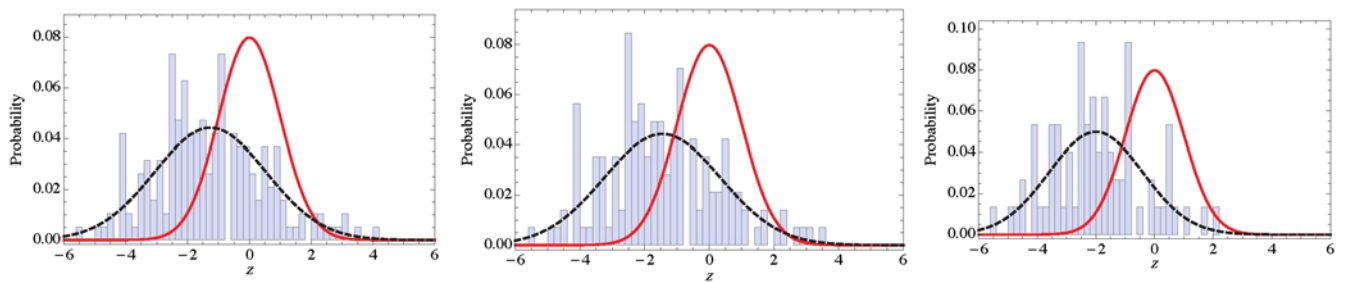


Figure 6. See the caption of Figure 5. In this case, the model tested is [Boore and Atkinson \(2008\)](#) at 2 Hz. From left to right: the maximum distances considered are 300, 200, and 100 km (corresponding to $\text{LLH} = 4.13, 4.48, 5.35$). The color version of this figure is available only in the electronic edition.

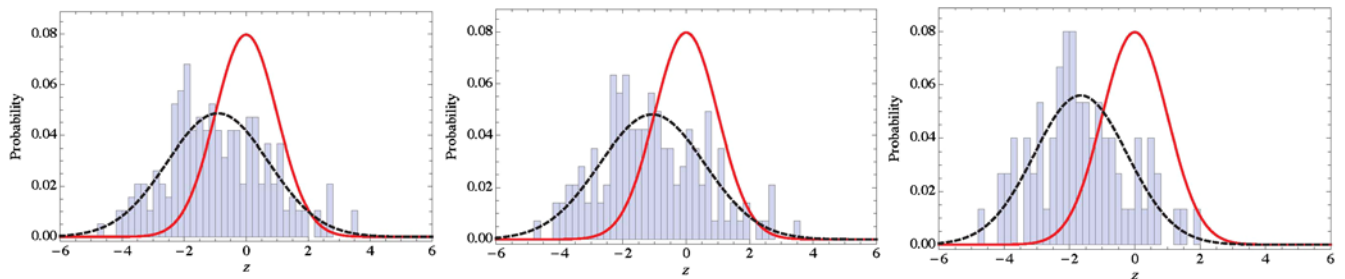


Figure 7. See caption of Figure 5. In this case, the model tested is [Zhao et al. \(2006\)](#) at 2 Hz. From left to right: the maximum distances considered are 300, 200, and 100 km (corresponding to $\text{LLH} = 3.4, 3.68, 4.3$). The color version of this figure is available only in the electronic edition.

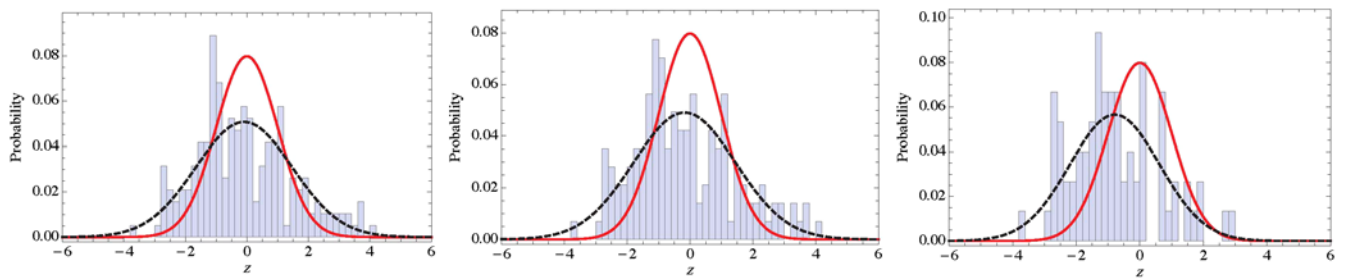


Figure 8. See caption of Figure 5. In this case, the model tested is [Akkar and Bommer \(2010\)](#) at 2.3 Hz. From left to right: the maximum distances considered are 300, 200, and 100 km (corresponding to $\text{LLH} = 2.68, 2.8, 2.76$). The color version of this figure is available only in the electronic edition.

Testing Predictions and Observations on the Japanese Data

The rich Japanese dataset contains both weak and strong ground motions, corresponding to a wide magnitude range. The KiK-net and K-NET network recordings have been collected up to the end of 2009 (Laurendeau *et al.*, 2011). Only events characterized in the F-net catalog are selected in order to have consistent metaparameters (M_w , hypocenter location, focal depth, and rake angle). Besides, the data subset used in this study includes only crustal events (focal depth ≤ 25 km and excluding offshore events on the subduction side) and rock sites ($V_{S30} \geq 500$ m/s, with V_{S30} deduced from KiK-net velocity profiles; Boore *et al.*, 2011). A magnitude–distance filter was applied according to Kanno *et al.* (2006) predictions, taking 2.5 Gal as a PGA threshold. *S*-wave triggered and multievent records have been eliminated. The source distance is the hypocentral distance for events with $M_w < 5.7$ and the closest distance from the fault plane to the observation site for events with larger magnitudes.

A set of eight models is selected, including recent models derived from Japanese data or from other active crustal regions of the world. First, the aim is to test a dataset with characteristics close to the French accelerometric dataset to determine if the number of recordings available is sufficient to consider the results reliable. Second, the same GMPEs are tested against the Japanese dataset in the moderate-to-large magnitude range. The objective is to analyze the performance of the models according to the magnitude range considered. In other words, using the Scherbaum *et al.* (2009) method, the ranking of the models obtained in the low-magnitude range is compared with the ranking of the models from the larger-magnitude range.

At first, only recordings corresponding to earthquakes with magnitudes between 4 and 4.9 are considered (with at least 8 recordings per event). The difference between this dataset and the French dataset is in the distance distribution: the Japanese network is much denser, and distances available in our database are shorter. For each run (each curve in Fig. 9), subsets are extracted at random from the initial dataset with a condition on the total number of recordings: the sample must contain 170 to 210 records (190 on average). The resulting sample is made up of 11 to 18 earthquakes, distributed all over Japan. An example of a random dataset is displayed in Figure 10. The range of LLH values obtained (up to 5.2) is comparable to the LLH values from the French weak-motion data. Three models provide low LLH, between 1.5 and 2.5, implying a good fit with the data (Kanno *et al.*, 2006; Cauzzi and Faccioli, 2008; Chiou and Youngs, 2008). The results confirm again that the fit between the predictions and the observations varies with the frequency considered. For frequencies greater than 6.0, all tested models provide stable LLH values over the frequency range, restricted to a narrow interval (LLH = 1.8–2.8). Most importantly for this test of stability, the results do not change much from one subset to another,

which implies that the rough hierarchy between models can be obtained with a small dataset of recordings. This dataset is the same size as the French weak-motion dataset tested but with a wider distance range in the case of the French dataset. The Cauzzi and Faccioli (2008), Kanno *et al.* (2006), and Chiou and Youngs (2008) models are identified as the best-fitting models, with the lowest LLH over the whole frequency range. The model with the poorest fit is Boore and Atkinson (2008). GMPEs have then been grouped according to their ranking (Table 3).

Next, all recordings corresponding to earthquakes with magnitudes between 5 and 7 are considered (Fig. 11, around 1200 recordings if considering events with at least 10 recordings). Again, LLH values obtained for the different GMPEs are in a rather narrow interval for frequencies from 6 to 10 Hz (1.9–2.4) but are quite different for frequencies less than 5–6 Hz (1.4–3.2). Three models emerge as the best-fitting equations between 0–10 Hz, with LLH values varying from 1.5 to 2 maximum: Cotton *et al.* (2008), Cauzzi and Faccioli (2008), and Kanno *et al.* (2006). It is worth noting that the Kanno *et al.* (2006) and Cotton *et al.* (2008) models have been derived from a database of Japanese recordings, whereas the Cauzzi and Faccioli (2008) model is based on a database made up of $\sim 80\%$ Japanese recordings. These results would tend to support the idea that ground motions in Japan display specific features (regional specificity). Some similar findings were obtained by Delavaud, Scherbaum, *et al.* (2012) based on two frequencies (1 Hz and PGA). Two other equations yield rather stable LLH values in the whole frequency range, Zhao *et al.* (2006) and Chiou and Youngs (2008), however, with slightly higher LLH values in the lower frequency range (< 4 Hz, $1.8 < \text{LLH} < 2.4$).

Based on the hierarchy obtained from the LLH values, the models are ranked in four categories (Table 3), from the best-fitting ($1.5 < \text{LLH} < 1.8$ for $f < 6$ Hz) to the worst ($1.7 < \text{LLH} < 3.4$ for $f < 6$ Hz) models. Comparing the results of the testing for low and larger magnitudes, it is interesting to observe that the hierarchy between models is only partly preserved. For two models, the ranking obtained from the low-magnitude dataset is quite different from the ranking obtained from the larger-magnitude dataset: the Cotton *et al.* (2008) model is no longer ranked among the best-fitting models (LLH around 2.5), whereas Chiou and Youngs (2008) is now among the best-fitting models (LLH around 2.0). Moreover, examples of normalized residuals are displayed in the case of the Cauzzi and Faccioli (2008) model (Fig. 12). The histograms highlight the link between a low LLH (~ 1.6) and the good fit of the normalized residual distribution with respect to the standard normal distribution.

Conclusions

We have analyzed and quantified the consistency between several GMPEs and three datasets, a low-magnitude (3.8–4.5) dataset of recordings from the French Accelerometric Network and two datasets built from the Japanese

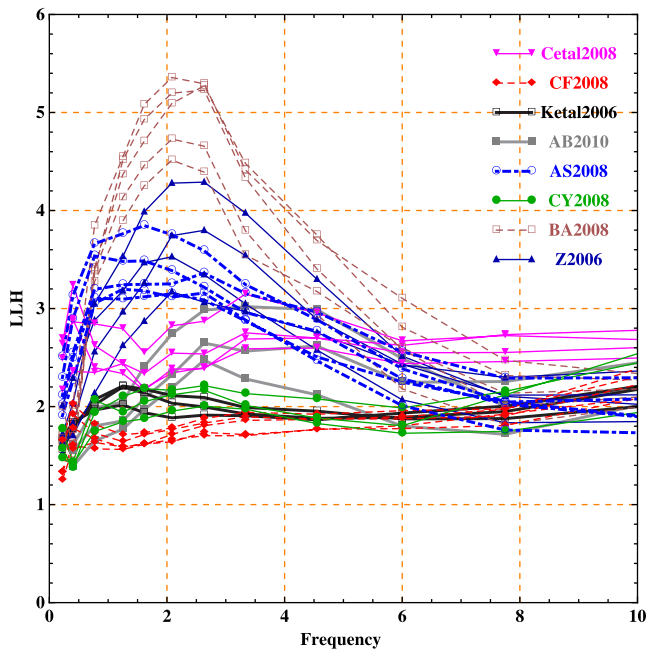


Figure 9. Testing GMPEs against the low-magnitude Japanese dataset (M 4–4.9): log-likelihood LLH values obtained on the Japanese dataset versus frequency. Five subsets are considered for each GMPE. See Table 2 for the GMPE abbreviations. Each subset is randomly extracted from the original dataset (condition: 170 to 210 recordings, resulting in 11 to 18 earthquakes). The color version of this figure is available only in the electronic edition.

K-NET and KiK-net networks. From these studies, we derive several conclusions and highlight the remaining key questions.

The Scherbaum *et al.* (2009) technique, relying on the calculation of a log-likelihood LLH, is a very practical and powerful tool to quantify the fit between predictive equations and observations. We find that for LLH values of 1.5–1.6, the distribution of the normalized residuals matches well a standard normal distribution; whereas for values greater than

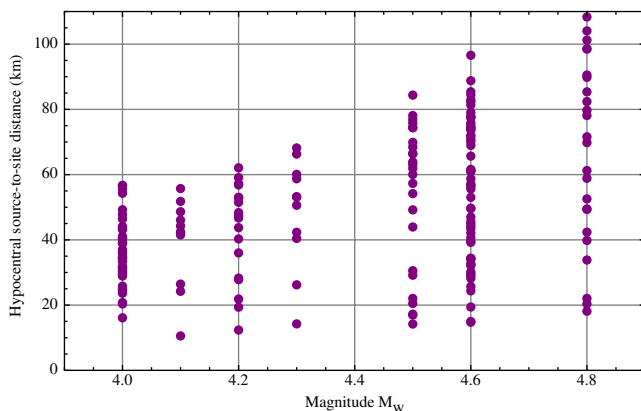


Figure 10. Example of a low-magnitude dataset randomly extracted from the original Japanese dataset (185 recordings and 11 events). The color version of this figure is available only in the electronic edition.

Table 3
Results of the Testing on the Japanese Dataset:
Classification of the GMPEs According to Their Fit
to the Data*

Rank of Models According to LLH	Larger Magnitude Range	Low Magnitude Range
Best fit	CF2008, Ketal06, Cetal08	CF2008, Ketal06, CY2008
Intermediate fit	Z2006, CY2008	Cetal08, AB2010
Poorly fit	AB2010, AS2008	AS2008, Z2006
Worst fit	BA2008	BA2008

*Two datasets have been considered separately (see the text).

~3–4, the mean, sigma, or both values calculated from the residual distribution strongly move away from the parameters of the standard normal distribution.

The fit between the observations and predictions proved in several cases to vary greatly with the frequency. When enough data are available, the testing and application of the Scherbaum *et al.* (2009) technique (or any other technique for testing GMPEs against data) should be carried out separately for each frequency. Otherwise, some information is lost, and a mean LLH is calculated that might not represent well the individual LLH for each frequency.

The analysis of the dataset from the French Accelerometric Network brings new insights for low-to-moderate seismicity regions of Western Europe (shallow active regions). The three models yielding the lowest LLH values on the French accelerometric low-magnitude dataset over the whole frequency range are the Cauzzi and Faccioli (2008), Akkar and Bommer (2010), and Abrahamson and Silva (2008) equations. These models, derived from different crustal tectonic environments, are thus, the equations most consistent with the weak-motion dataset recorded in active regions of France (Alps, Pyrenees, and Lower Rhine Embayment). Both models CF2008 and AB2010 have been selected in the SHARE project for application in active shallow crustal

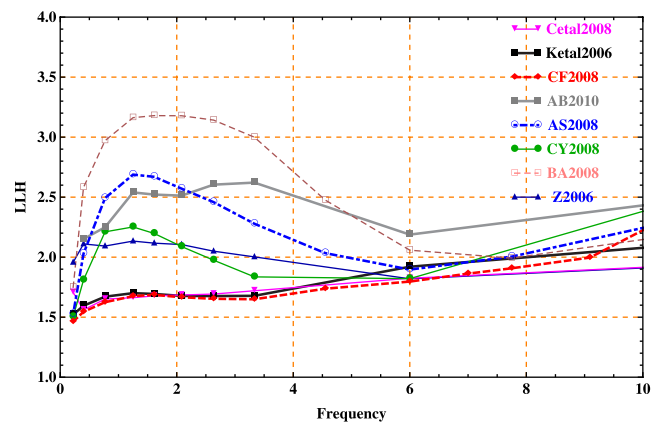


Figure 11. Testing GMPEs against the moderate-to-large magnitude Japanese dataset (M 5–7): log-likelihood LLH values versus frequency. See Table 2 for the GMPE abbreviations. The color version of this figure is available only in the electronic edition.

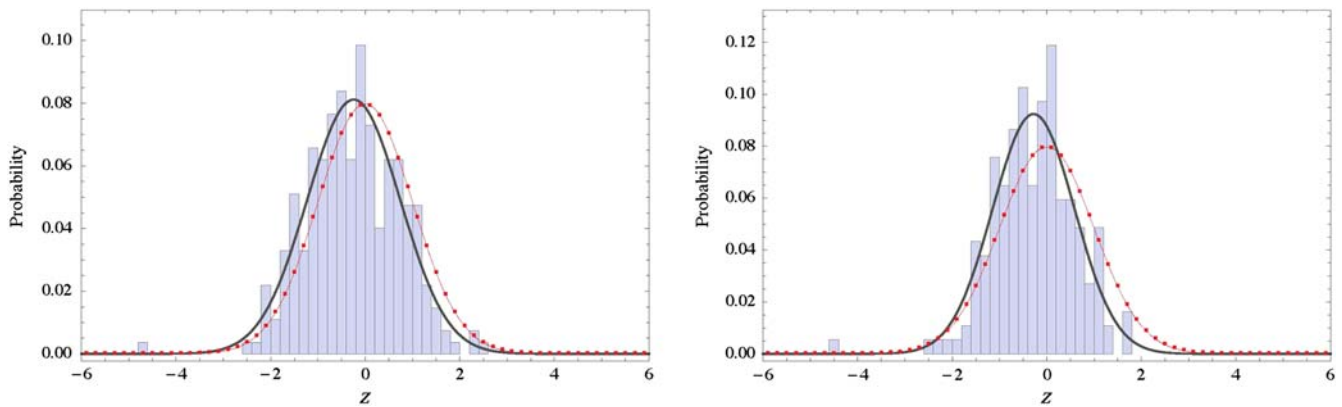


Figure 12. Predictions from the [Cauzzi and Faccioli \(2008\)](#) model compared to the Japanese accelerometric subsets (at 1.25 Hz): histogram of the normalized residuals superimposed on the standard normal distribution representing the model. Left: subset containing 1143 recordings corresponding to 36 events with $5 \leq M_w \leq 7$ (LLH = 1.65). Right: subset containing 185 recordings corresponding to 15 events with $4 \leq M_w \leq 4.9$ (LLH = 1.6). The Gaussian with mean and sigma calculated from the residuals is superimposed on the histogram. The color version of this figure is available only in the electronic edition.

regions across Europe. [Akkar and Bommer \(2010\)](#) is a pan-European model, whereas [Cauzzi and Faccioli \(2008\)](#) mostly relies on Japanese data, and [Abrahamson and Silva \(2008\)](#) on Californian and worldwide data. This result does not highlight the regional variation of ground motions.

The sensitivity studies carried out on the Japanese database show that, considering a subset with properties similar to the French accelerometric dataset (in terms of magnitudes and amount of recordings), the ranking of GMPEs obtained does not depend on the subset. This finding implies that the results obtained from the French accelerometric dataset, made up of 191 recordings at distances less than 300 km, can be considered to be stable (until more data are available to prove it, especially more short-distance recordings).

The same set of predictive models is tested on the moderate-to-large magnitude dataset from Japan ($5 \leq M \leq 7$). Comparing the ranking of GMPEs obtained from this large-magnitude range with the ranking resulting from the low magnitudes, some features are maintained, but for some models the ranking is modified. The magnitude scaling is therefore controlling the ground motions for some of the GMPEs tested. An interesting observation is that, when considering magnitudes that fall in the validity limits of the equations, the models predicting the best observations are all native Japanese models (with LLH values of 1.5–1.7, indicating that the fit is very good). Ground-motion scaling in Japan might differ significantly from other active regions.

Data and Resources

All of the accelerometric data from the French Accelerometric Network (<http://www-rap.obs.ujf-grenoble.fr/>, last accessed April 2012) as well as the accelerometric data from Japan (www.k-net.bosai.go.jp, last accessed April 2012) and the earthquake data from the Réseau National de Surveillance Sismique (RéNaSS, <http://renass.u-strasbg.fr/>, last

accessed April 2012) are available online. The SHARE project is presented at <http://www.share-eu.org/> (last accessed April 2012).

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Swiss Seismological Service
Institute of Geophysics, ETH Zurich, NO
Sonnegstrasse 5
8092 Zurich, Switzerland
(E.D.)

Institute of Earth and Environmental Sciences
University of Potsdam
Karl-Liebknecht-Strasse 24-25
14476 Golm, Germany
(N.K.)

ISTerre
IRD, Université Joseph Fourier, CNRS, IFSTTAR
BP 53
38041 Grenoble Cedex 9, France
Celine.Beauval@obs.ujf-grenoble.fr
(C.B., H.T., A.L., F.C., P.G.)

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